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# Renegotiating Home Mortgages: Evidence from the Subprime Crisis

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# RENEGOTIATING HOME MORTGAGES: EVIDENCE FROM THE SUBPRIME CRISIS\*

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## Abstract

We document that servicers have been reluctant to renegotiate mortgages since the foreclosure crisis started in 2007, performing payment reducing modifications on about 3 percent of seriously delinquent loans. We show that this reluctance does not result from securitization: servicers renegotiate similarly small fractions of loans that they hold in their portfolio. Our results are robust to different definitions of renegotiation, including the one most likely to be affected by securitization, and different definitions of delinquency. Our results are strongest in subsamples in which unobserved heterogeneity between portfolio and securitized loans is likely to be small and for subprime loans. We use a theoretical model to show that *redefault risk*, the possibility that a borrower still defaults despite costly renegotiation, and *self-cure risk*, the possibility that a seriously delinquent borrower will become current without renegotiation, make renegotiation unattractive to investors.

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# 1 Introduction

Many commentators have attributed the severity of the foreclosure crisis in the U.S. in the 2007-2009 period to the unwillingness of lenders to renegotiate mortgages, and, as a consequence, have placed renegotiation at the heart of the policy debate. Every major policy action to date has involved encouraging lenders, in one way or another, to renegotiate loan terms in order to reduce borrower debt loads. According to the Treasury-sponsored HOPENow initiative in December of 2007, lenders were expected to prevent adjustable rate mortgages from increasing to higher rates at the first reset of the mortgage.<sup>1</sup> “Hope For Homeowners”, enacted by Congress in July of 2008, envisioned that lenders would write-off a substantial portion of the principal balance of mortgages for financially distressed households.<sup>2</sup> The Obama Administration’s Making Home Affordable Plan, announced in February of 2009, provided financial incentives to servicers to renegotiate loans on the condition that the lenders reduced the interest rate for a significant period of time.<sup>3</sup>

The appeal of renegotiation to policy makers is simple to understand. If a lender makes a concession to a borrower by, for example, reducing the principal balance on the loan, it can prevent a foreclosure. This is clearly a good outcome for the borrower, and possibly good for society as well. But the key to the appeal of renegotiation is the belief that it can also benefit the lender, as the lender only loses money if the reduction in the value of the loan exceeds the loss the lender would sustain in a foreclosure. In short, according to its proponents, renegotiation of home mortgages is a type of public policy holy grail, in that it helps both borrowers and lenders at little or no cost to the government.<sup>4</sup>

In this paper, we explore the renegotiation of home mortgages using data from LPS (Lender Processing Services), a large, detailed sample of residential mortgages. Our primary empirical analysis involves following borrowers over the year subsequent to their first serious delinquency, and counting the frequency of renegotiation.<sup>5</sup> Measuring renegotiation in the LPS data is a challenge because there is no field in the data that identifies whether or not a servicer has changed the terms of, or “modified”, the loan. We overcome this difficulty by developing an algorithm to identify modifications that we validate on an unrelated dataset which includes a modification flag.

We explore several different definitions of renegotiation in the data. Our first measure is the number of concessionary modifications, which we define as either reductions in the

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<sup>1</sup>Edmund L. Andrews, In Mortgage Plan, Lenders Set Terms, N.Y. TIMES, Dec. 7, 2007.

<sup>2</sup>“Bush Signs Wide-Ranging Housing Bill Into Law,” *Wall Street Journal*, July 31, 2008.

<sup>3</sup>See “\$275 Billion Plan Seeks To Address Crisis In Housing,” *New York Times*, Feb. 18, 2009.

<sup>4</sup>See this discussion in Congressional Oversight Panel (2009), Zingales (2008) and Geanakoplos and Koniak (2008), as examples.

<sup>5</sup>Until 2008, the dataset was known as McDash.

interest rate, reductions in the principal balance, extensions of the term, or combinations of all three, which serve to reduce a borrower's monthly payment. This is a key focus of our analysis because there is a consensus among many market observers that concessionary modifications are the most, or possibly the only, effective way of preventing foreclosures. As the Congressional Oversight Panel wrote, "Any foreclosure mitigation plan must be based on a method of modifying or refinancing distressed mortgages into affordable ones. Clear and sustainable affordability targets achieved through interest rate reductions, principal write-downs, and/or term extensions should be a central component of foreclosure mitigation."<sup>6</sup>

Because the Pooling and Servicing Agreements (PSAs) which govern the conduct of servicers when loans are securitized often place limits on the number of modifications a servicer can do we broaden our definition of renegotiation to include any modification, regardless of whether it lowers the borrower's payment. Modifications are often thought to always involve concessions to the borrower but many, and in some subsets, most modifications involve the capitalization of arrears into the balance of the loan and thus lead to increased payments.

Finally, we attempt to include in our measure of renegotiation the number of times that lenders allow borrowers to extinguish their liabilities by repaying less than the outstanding balance of the loan, transactions known as short-payoffs, short-sales or deed-in-lieu of foreclosures, depending on the structure. We do this by counting the number of seriously delinquent loans that the servicer reports as "paid off."

No matter which measure of renegotiation we use, one message is quite clear: lenders rarely renegotiate. Less than 3 percent of the seriously delinquent borrowers in our sample received a concessionary modification in the year after the first serious delinquency. More borrowers received modifications using our broader definition, but the total still accounted for less than 8 percent of the seriously delinquent borrowers. And finally, less than 5 percent of all of our troubled borrowers repaid their mortgages, putting an upper-bound on the number who could have repaid less than the principal balance of the loan. These numbers are small both in absolute terms as well as relative to approximately half of the sample for which foreclosure proceedings were initiated and nearly 30 percent for which they were also completed.

We then turn to the question of why renegotiation is so rare. If the logic described in the second paragraph is right, lenders should find renegotiation attractive, even in the absence of government prodding. Yet, we observe very little renegotiation in the data. We address this apparent paradox.

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<sup>6</sup>See the Congressional Oversight Panel (2009). This view is widely held and the main focus of the administration Making Home Affordable foreclosure prevention plan was to encourage servicers to modify loans to reduce monthly payments to 31% of income.

The leading explanation attributes the reluctance of lenders to renegotiate to the process of securitization.

The complex webs that securitization weaves can be a trap and leave no one, not even those who own the loans, able effectively to save borrowers from foreclosure. With the loan sliced and tranced into so many separate interests, the different claimants with their antagonistic rights may find it difficult to provide borrowers with the necessary loan modifications, whether they want to ‘or not. In the tranche warfare of securitization, unnecessary foreclosures are the collateral damage. (Eggert (2007))

More precise institutional evidence appears to confirm the role of securitization in hobbling renegotiation. As mentioned above, PSAs do often place global limits the number of modifications a servicer can do for a particular pool. The rules by which servicers are reimbursed for expenses may provide a perverse incentive to foreclose rather than modify. And because servicers don’t internalize the losses on a securitized loan, they may not behave optimally. Another issue is the possibility that those investors whose claims are adversely affected by modification will sue. Finally, historically SEC rules have said that contacting a borrower who is less than 60 days delinquent constitutes an ongoing relationship with the borrower and jeopardizes the off-balance sheet status of the loan.

But some market observers express doubts about the renegotiation-limiting role of securitization. Hunt (2009) conducted an exhaustive review of a sample of PSA’s and conclude that, “it appears that large-scale modification programs may be undertaken without violating the plain terms of PSAs in most cases.” Although some servicers have expressed concern about lawsuits, of the more than 800 lawsuits filed by investors in subprime mortgages through the end of 2008, not one involved the right of a servicer to modify a loan.<sup>7</sup> Even the Congressional Oversight Panel (2009), which did view securitization as a problem in general, conceded that, “The specific dynamics of servicer incentives are not well understood.” Finally, the SEC ruled in 2008 that if default was “reasonably foreseeable” then contact with a borrower prior to 60-day delinquency would not affect the accounting status of the loan.

Our empirical analysis provides strong evidence against the role of securitization in preventing renegotiation. The LPS data includes loans that are serviced for private securitization trusts that are not sponsored by any of the government sponsored entities, so-called “private label” loans, which are subject to all the contract problems described above. It also includes loans that are owned by servicers, so-called “portfolio” loans, which are immune

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<sup>7</sup>Navigant report, Congressional Oversight Panel (2009).

to such problems. We compare renegotiation rates controlling for observable characteristics of the loans. For our narrowest definition of renegotiation, payment-reducing modification, we find that the differences in the likelihood of renegotiation in the twelve months subsequent to the first 60-day delinquency between the two types of loans is neither economically nor statistically significant. When we consider the broader definitions that includes any modification at all, which, as we mentioned above, we would expect to be most affected by securitization, the data even more strongly reject the role of securitization in preventing renegotiation. We also find that servicers are *more* likely to perform modifications, broadly defined, and to allow the borrower to prepay on a private-label loan compared to a portfolio loan.

Our results are highly robust. One potential problem with the data is that there is unobserved heterogeneity in the characteristics of portfolio and private-label loans. To address this, we exploit subsets of the LPS data, in which servicers provide an exceptional amount of information about borrowers. When we exclude observations where the servicer failed to report whether the borrower fully documented income at origination or what the debt-to-income ratio was at origination, our results become even stronger. When we focus on only loans for which the borrower fully documented income, we obtain results that are broadly consistent or, in some cases, stronger than the results for the full sample. Finally, we limit our sample to only subprime loans (as defined in LPS). These loans comprise only 7 percent of the LPS data, but account for more than 40 percent of all serious delinquencies and almost 50 percent of the modifications that we identify in the data. The results that we obtain for the subprime sample are also consistent with our results for the full sample.

Another potential issue with our focus on 60-day delinquent loans is that portfolio lenders can contact borrowers at any time, whereas some securitization agreements forbid lenders from contacting borrowers until they are at least sixty days delinquent (two missed payments). When we shift our focus to 30-day delinquent borrowers (one missed payment), our results continue to show no meaningful difference between renegotiation of private-label and portfolio loans.

One other possibility is that our algorithm for identifying modifications is somehow missing a class of loan-mitigation actions taken by servicers. Forbearance agreements and repayment plans, for example, would not necessarily show up in our data. However, neither of these actions constitute renegotiation in any classic sense, because the lender still expects the borrower to repay in full, including interest on any delayed payment. In addition, unlike modifications, PSAs never place any limits on the use of forbearance agreements or repayment plans, so *a priori*, we would have less reason to expect a difference in their use across private-label and portfolio loans. Finally, most successful forbearance agreements conclude

with a modification to allow the borrower to repay the arrears incurred in forbearance. With all of that said, we test the proposition that servicers engage in other loss mitigation actions by looking at the “cure rate.” This is the percentage of loans that transition to current status after becoming sixty days delinquent. We find that in the full sample, private-label loans are less likely to cure, but that the gap, although statistically significant, is small – correcting for observable characteristics, we estimate a cure rate of around 30 percent for the typical portfolio loan and a cure rate of about 2 percentage points less for an otherwise equivalent private-label loan. However, for the subprime subsample, the subsample with information about documentation status, and the sample of fully documented loans, we find that private-label loans are significantly *more* likely to cure.

So if contract frictions are not a significant problem, then why don’t lenders renegotiate with delinquent borrowers more often? We argue for a very mundane explanation: lenders expect to recover more from foreclosure than from a modified loan. This may seem surprising given the large losses lenders typically incur in foreclosure, which include both the difference between the value of the loan and the collateral and the substantial legal expenses associated with the conveyance. The problem is that renegotiation exposes lenders to two types of risks which can dramatically increase its cost. The first is what we will call “self-cure” risk. As we mentioned above, more than 30 percent of seriously delinquent borrowers “cure” without receiving a modification, which, if taken at face value, means that, in expectations, thirty percent of the money spent on a given modification is wasted. The second cost comes from borrowers who re-default; our results show that a large fraction of borrowers who receive modifications end up back in serious delinquency within six months. For them, the lender has simply postponed foreclosure; in a world with rapidly falling house prices, the lender will now recover even less in foreclosure. In addition, a borrower who faces a high likelihood of eventually losing the home will do little or nothing to maintain the house, or may even contribute to its deterioration, again reducing the expected recovery by the lender.

In Section 4 of the paper, we formalize the basic intuition of the investor renegotiation decision, with a simple model. We show that the higher cure rates, higher re-default rates, higher expectations of house price depreciation, and a higher discount rate all make renegotiation less attractive to the investor. Thus one cannot evaluate a modification by comparing the reduction in the interest rate on the loan or in the principal balance with the expected loss in foreclosure. One must take into account both the redefault *and* self-cure risks, something that most proponents of modification do not do.<sup>8</sup>

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<sup>8</sup>Many proponents of aggressive modification take into account redefault risk, and the MHA plan did address it by providing some insurance to investors who modified loans for further house price declines.

To our knowledge, this paper is the first to directly estimate the likelihood of renegotiation of private label and portfolio-held loans. Piskorski, Seru, and Vig (2009) address the question of the effects of securitization on renegotiation, but rather than directly identify renegotiation, they run “black-box” foreclosure regressions using LPS data and argue that observed differences in foreclosure rates *imply* differences in renegotiation activity. Our results contradict this interpretation. For renegotiation to explain the differences in foreclosure rates, there would have to be large errors in our algorithm for identifying renegotiation, and those errors would have to be significantly biased toward portfolio loans, a possibility that is particularly problematic given that the type of renegotiations we focus on are precisely the ones that PSAs supposedly prevent. In addition, most of the loan histories in the LPS sample are right-censored, meaning that the borrowers have neither lost their homes nor paid off their mortgages when the data end, making it impossible to equate the absence of a foreclosure with successful renegotiation. By contrast, a “cure” is a necessary condition for renegotiation, and thus the differences we report in cure rates across portfolio and private-label loans, which are neither large, nor of consistent sign, contradict the claim that securitization is a major obstacle to renegotiation.

The implications of our research for policy are three-fold. First, “safe harbor” provisions which shelter servicers from investor lawsuits are unlikely to affect the number of modifications should have little effect. Second, and more broadly, the number of “preventable foreclosures” may be far less than many believe.

Finally, we point out that while our model shows why investors may not want to do modifications, that does not necessarily imply that modifications may not be socially optimal. One key input in our theoretical model is the discount rate and it is possible that investors, especially in a time when liquidity is highly valued, may be less patient than society as a whole and therefore foreclose when society would prefer renegotiation. Large financial incentives to investors or even to borrowers to continue payment could mitigate this.

## 1.1 Related Literature and Existing Evidence

Our research draws on existing literature in several different fields. First, there has been substantial interest in the question renegotiation of home mortgages among real estate economists both prior to and as a result of the current crisis. Riddiough and Wyatt (1994a), Riddiough and Wyatt (1994b) and Ambrose and Capone (1996) addressed information issues

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However, none of the proponents ever mention self cure risk even though it is well-known in the industry, see: <http://www.calculatedriskblog.com/2008/09/loan-modifications-anecdotes-and-data.html>.

that inhibit efficient renegotiation which we draw on extensively in Section 4. Springer and Waller (1993), in a early example, explores patterns in the use of forbearance as a loss mitigation tool. Capone (1996) and Cutts and Green (2005) both discuss the institutions, with the providing historical evidence and focusing on issues in the mid-1990s and the latter discussing innovations since then.

The start of the subprime crisis in 2007 led to a resurgence of interest in the topic among real estate economists and new interest from other field, in particular, law. In real estate, Quercia, Ding, and Ratcliffe (2009), Cutts and Merrill (2008), Stegman, Quercia, Ratcliffe, Ding, Davis, Li, Ernst, Aurand, and Van Zandt (2007) and Mason (2007), all discuss issues with contemporary loss mitigation approaches. Legal researchers, see White (2008) and White (2009), for example, have addressed empirical questions about the frequency and characteristics of loan modifications, closely related to the analysis in this paper. In addition, they have also looked at issues related to the restrictions imposed by contracts (Hunt (2009), Gelpern and Levitin (2009)) and the interaction between foreclosure, renegotiation and personal bankruptcy (Levitin (2009b), Levitin (2009a)).

More broadly, real estate economists have explored the factors that lead delinquent mortgages to transition to foreclosure or to cure, one of which is renegotiation. Pre-crisis papers include Ambrose and Capone (1998), Ambrose, Buttimer Jr, and Capone (1997), Ambrose and Capone (2000), Lauria, Baxter, and Bordelon (2004), Danis and Pennington-Cross (2005), Pennington-Cross (2009) and Pennington-Cross and Ho (2006). Mulherin and Muller (1987) discusses conflicts between mortgage insurers and owners that may lead servicers to induce or postpone foreclosure inefficiently. In the light of the crisis, Piskorski, Seru, and Vig (2009) Cordell, Dynan, Lehnert, Liang, and Mauskopf (2008a) have revisited the question.

The issue of dispersed ownership and debt renegotiation has received a fair amount of attention in the corporate finance literature. Gan and Mayer (2006), for example, focus on commercial mortgages and find that servicers delay liquidation of delinquent mortgages when they are also the holders of the equity tranche of the deal. This suggests that participating in the losses due to liquidation may alleviate some of the agency problems posed by the separation of ownership and servicing pointed out before. However, it may also lead to conflicts of interest between holders of different tranches. In their setting, Gan and Mayer (2006) find that the servicers' behavior is consistent with asset substitution, as servicers seek to benefit from the option-like payoff of their position. Also, the contractual restrictions imposed by pooling and servicing agreements (discussed above) and standard economic arguments on the effects of dispersed ownership of debt (as in Bolton and Scharfstein 1996 and Asquith, Gertner, and Scharfstein 1994) further reduce the incentives of servicers to

modify mortgages.

## 2 Data

We use a dataset constructed by Lender Processing Services (hereafter referred to as LPS). This is a loan-level dataset that covers approximately 60 percent of the U.S. mortgage market, and contains detailed information on the characteristics of both purchase-money mortgages and mortgages used to refinance existing debt. This dataset is especially useful in the context of this paper, as it includes both securitized mortgages and loans held in portfolio.<sup>9</sup> The LPS data specifically denote whether a mortgage is held in portfolio, or securitized by a non-agency, private institution.<sup>10</sup> If institutional constraints are restricting the modification process for private-label, securitized loans, we would expect to see relatively few modifications among them, as compared to portfolio loans. Unfortunately, our LPS sample does not include direct information regarding loan modifications.<sup>11</sup> However, LPS does provide monthly updates to loan terms, so it is possible to identify loan modifications indirectly (and imperfectly). Table 1 shows two examples of modifications in the data. In the first example, the servicer cuts the interest rate, capitalizes arrears into the balance of the loan, and extends the term of the loan to 40 years. In the second example, the servicer just capitalizes arrears into the balance of the loan. In both cases the loan is reported as “current” after the modification, whereas before it was 90+ days delinquent.

We denote a loan as being modified if there is a change in its terms that was not stipulated by the initial terms of the contract. Such modifications include interest-rate reductions, principal-balance reductions, and term extensions. We can also identify principal-balance and mortgage-payment *increases* that reflect the addition of arrears into the balance of a loan.<sup>12</sup> We spell out our algorithm for identifying modifications in more detail in Appendix

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<sup>9</sup>For a more detailed discussion of the LPS data, we direct the reader to Foote, Gerardi, Goette, and Willen (2009).

<sup>10</sup>The LPS data also denote when a loan is securitized by a GSE (Government Sponsored Enterprise) such as Freddie Mac or Fannie Mae. We eliminate this class of loans since the GSEs hold all credit risk, and thus are not subject to any modification restrictions.

<sup>11</sup>In a recent report, the Office of Thrift Supervision (OTS), in collaboration with the Office of the Comptroller of Currency (OCC), used data from LPS to analyze the outcomes of recent mortgage modification programs (OCC and OTS Mortgage Metrics Report, Third Quarter 2008) In this report, they had access to supplementary data from servicers that included the identification of loans in the LPS data that had been modified. We have not been able to obtain access to this data.

<sup>12</sup>One of the major types of loan modifications that we are largely unable to identify are interest rate freezes for subprime ARMs, which reset after two or three years. However, the reason that we cannot identify those freezes is because many are not binding; the fully-indexed rate is lower than the initial rate. These modifications will have no major effect on the current terms of the mortgage, so we do not view this as a major drawback.

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There are two potential mistakes we can make in this exercise. First, we may falsely identify modifications (“false positives”) because of measurement error in the data (for example, a mistake in the updated balance or interest rate), or some endogenous behavior on the part of the borrower (for example, a borrower making extra principal payments). Second, we could miss modifications (“false negatives”) because our algorithm for finding modifications is incomplete. In order to test our algorithm we use data from the Columbia files put together by Wells Fargo’s CTSLink service. This data includes a similar set of variables as the LPS dataset (on performance of the loans and characteristics of the borrower at origination), but is limited to private-label loans. These files include, however, explicit flags for modifications. This allows us to use the same algorithm described in Appendix A, and compare the modifications we identify to the “true” modifications. Results are reported in Table 2. Overall our algorithm performs well, with 17% false negatives (i.e. we do not identify around 17% of the “true” modifications) and around the same percentage of false positives (i.e. approximately 17% of the modifications we identify are not flagged as modifications on the CTSLink data). By type of modification, our algorithm performs best for principal reductions, term increases and fixed-rate mortgage reductions, whereas it performs comparatively worse for ARM rate reductions and for principal increases.

## 2.1 Summary Statistics from the Data

Table 3 reports the number of modifications performed each quarter from the first quarter of 2007 through the final quarter of 2008, disaggregated by the type of modification. Each of the numbers is a multiple of 10 because we used a 10% random sample and scaled up the numbers we found. The first column of Table 3 simply reports the total number of loan modifications made. Not surprisingly, modifications have become more common as the housing market has weakened. There appear to be more than 7-8 times as many modifications performed in the fourth quarter of 2008 as compared to the first quarter of 2007. In addition to the rapid growth in loan modifications, the composition of modifications has changed over time. This can be seen in the remaining columns of Table 3, which list the incidence of modifications of different types.<sup>13</sup>

An interesting finding is that most modifications entailed *increases* in the principal balance of a mortgage. Such increases are likely due to the addition of arrears to the outstand-

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<sup>13</sup>In many cases a mortgage will experience multiple types of modifications at the same time. For example, we see cases in the data in which the interest rate is decreased and at the same time the term of the loan is extended. Thus, the percentages in Table 3 are not calculated with respect to the number of loans modified, but rather with respect to the number of modifications performed.

ing mortgage balance for delinquent borrowers, and often increase the monthly mortgage payment by a nontrivial amount. While the absolute numbers of balance-increasing modifications are still rising, they are falling as a percentage of total modifications. In the last few quarters, interest-rate reductions, which necessarily involve a decrease in the monthly mortgage payment, have become more frequent, rising to more than 26 percent of all modifications performed in 2008:Q4. Table 3 provides further information regarding the behavior of monthly mortgage payments for loans that have undergone a modification. There are several notable patterns in this table. First, as of 2008:Q4, modifications that involved payment decreases were more common than those that involved payment increases. Furthermore, the average and median magnitude of payment decreases has recently increased in our sample. From 2007:Q1 to 2008:Q2 the average payment decrease ranged from approximately 10% to 14%, but then increased to approximately 20% and 22% in 2008:Q3 and 2008:Q4, respectively. Based on the logic from our simple framework above, it is likely that these will have more success than modifications involving increases in the payment and/or balance.

Another interesting observation from Table 3 is that the incidence of principal reductions is quite low in our data. This is likely due to two factors. First, the LPS data underrepresents the subprime mortgage market.<sup>14</sup> A few servicers that focus almost exclusively on subprime mortgages have recently begun modification programs that involve principal reduction.<sup>15</sup> In addition, from a theoretical perspective, principal reduction plans suffer from the severe incomplete-information problem noted earlier. Balance reductions are appealing to both borrowers in danger of default and those who are not. In a recent paper, we argued that to avoid such moral hazard concerns, lenders have a strong incentive to only provide modifications to those borrowers who are most likely to default.<sup>16</sup> Table 3 contains summary statistics regarding the characteristics at origination of both the sample of modified mortgages and the sample of all loans in the LPS dataset. The patterns that emerge from the table are consistent with such an argument. We discuss this point in more detail below. The sample of modified mortgages is characterized by substantially lower credit scores, higher ltv ratios, and slightly higher debt-to-income ratios. The discrepancy in ltv ratios may be underestimated, as the percentage of mortgages with an ltv ratio of exactly 80% is significantly higher in the modification sample compared to the full sample. As we argued

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<sup>14</sup>The majority of subprime mortgages are securitized by non-agency firms, and the LPS data includes approximately 35 percent of mortgages securitized by non-agency corporations.

<sup>15</sup>According to an October report by Credit Suisse, Ocwen Loan Servicing, LLC and Litton Loan Servicing LP were the only subprime servicers that had performed a nontrivial number of principal reduction modifications. Both of these servicers do not contribute to the LPS dataset.

<sup>16</sup>See Foote, Gerardi, and Willen (2008) for a more detailed discussion.

above, this likely implies a larger fraction of highly leveraged loans, for which the second liens are not observable in the data. In addition, the modification sample includes a higher fraction of mortgages with non-traditional amortization schedules, such as interest-only loans, option ARMs, hybrid-ARMs, and subprime loans.

In Table 4 we compare the size of payment decrease and payment increase modifications for loans held in a private-label trusts and loans held in portfolio. The results are somewhat mixed, as the size (as a percentage of the original payment) of the median payment decrease modification is larger for private-label loans in the first three quarters of 2008, but is smaller in the final quarter. We see a similar pattern for the median payment increase modification, while the differences are small for the mean and median payment increase.

### 3 Differences in Modification Behavior

In this section we directly address the question of whether the incidence of modification is impeded by the process of securitization. We show evidence that private-label loans and portfolio loans perform similarly, both unconditionally and when observable differences between securitized and portfolio-held loans are controlled for, using both a logit model with a 12-month horizon, as well as a Cox proportional hazard model that takes into account the problem of right censoring in the data.

In this section we will use several different definitions of modification, to make sure that our results are robust to the type of modification performed. Our first measure is the number of concessionary modifications, which we define as either reductions in the interest rate, reductions in the principal balance, extensions of the term, or combinations of all three, which serve to reduce a borrowers monthly mortgage payment. We use this as our primary definition of modification in our analysis, as there is a consensus among most market observers that concessionary modifications are the most, or perhaps the only, effective way of preventing foreclosures. Because Pooling and Servicing Agreements, which govern the conduct of servicers when loans are securitized, often limit modifications that change *any* of the contract terms (not just those that result in payment decreases), we broaden our definition of renegotiation to include any modification, regardless of whether it lowers the borrowers payment. As we discussed above, many, and in some subsets, most modifications, involve the capitalization of arrears into the balance of the loan, and thus lead to increased payments. Finally, we attempt to include in our measure of renegotiation the number of times that lenders allow borrowers to extinguish their liabilities by repaying less than the outstanding balance of the loan. These transactions are known as short-payoffs; short-sales or deed-in-lieu of foreclosures, depending on the structure. We do this by counting the

number of seriously delinquent loans that the servicer reports as paid off, and including these observations in our definition of modification.

Before turning to the regressions however, it is instructive to look at the unconditional frequencies of modifications in the data. Panel A of Table 5 shows the unconditional frequencies for each type of investor. The first takeaway from the table is the extremely low percentages of modifications for *both* types of mortgages. Only 3 percent of 60-day delinquent loans received concessionary modifications in the 12 months following the first serious delinquency, and only 8.5% of the delinquent loans received *any* type of modification in the same period. These are extremely low levels of modifications, and suggests that even if there are contract frictions that are preventing modifications in securitized trusts, the economic effects are small. The second takeaway from the table is the unconditional differences between portfolio loans and private-label loans are very small in absolute terms. There is a difference of approximately 0.6 percentage points and 0.3 percentage points for concessionary modifications and all modifications, respectively. These are very small differences, and suggest that contract frictions do not play an important role in inhibiting the renegotiation process for loans in securitized trusts. However, these are unconditional statistics, and it is possible that once observable differences in the characteristics of each type of loan and borrower are accounted for, the results may change.<sup>17</sup> Thus, we now estimate differences in modification behavior while controlling for observable loan and borrower characteristics. These characteristics include the contract interest rate at origination, the credit score of the borrower at origination, the loan-to-value ratio of the mortgage (not including second or third liens) at origination<sup>18</sup>, the logarithm of the nominal dollar amount of loan, whether the purpose of the loan was a refinance of a previous mortgage or a home purchase, whether the loan was considered to be subprime<sup>19</sup>, a measure of the amount of equity in the property at the time of delinquency, specified as a percentage of the original loan balance, and updated by state-level house price indexes, calculated by the Federal Housing Finance Agency (FHFA),<sup>20</sup> (and an indicator for a borrower who is in a position of negative equity at the

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<sup>17</sup>For example, if private-label loans are significantly riskier, and thus better candidates for modification on average, then the unconditional difference will significantly understate things.

<sup>18</sup>Because of the lack of information on second liens in the LPS data, and the prevalence of second mortgages as a way to avoid paying mortgage insurance, we include an indicator variable if the ltv ratio is exactly equal to 80 percent. These are the borrowers who likely took out second mortgages, as the requirement for mortgage insurance occurs at ltv ratios above 80 percent. Our experience with other, more complete datasets also confirms that many of these borrowers are likely to have second mortgages that bring the cumulative ltv ratio up to 100 percent.

<sup>19</sup>This definition of subprime comes from the mortgage servicers that contribute to the LPS dataset.

<sup>20</sup>House prices are measured at the state-level using the FHFA index. We also tried using Case-Shiller metropolitan area house price indexes, and found no substantive differences. We chose to use the OFHEO prices for our primary specifications due to their greater sample coverage.

time of delinquency, where the value of the mortgage exceeds the value of the home), and the unemployment rate of the county in which the borrower resides, calculated by the Bureau of Labor Services (BLS).<sup>21</sup> We also include, but do not show for space considerations, a set of cohort dummies that control for the quarter in which the mortgage was originated, information regarding the amortization schedule of the mortgage (interest-only or negative amortization, which include mortgages commonly referred to as option-ARMs), an indicator for whether the size of the mortgage is greater than the GSE conforming loan limits, an indicator for whether the house is a primary residence, an indicator for adjustable rate mortgages that contain a reset provision (so-called “hybrid ARMs”), and finally an indicator for a borrower that does not use the corresponding property as a principal residence (this includes both properties used strictly for investment purposes, as well as vacation homes).

### 3.1 Canonical Specification Results

Panel B of Table 5 displays the estimated marginal effects from a set of logit models for the three different types of modification definitions. The dependent variable is 1 if a 60-day delinquent loan is modified at any point in the 12 months following the first delinquency. The first column considers payment reducing (concessionary) modifications, the second column includes both payment reducing and payment increasing modifications, and the third column contains all modifications considered before, as well as prepayments. In all regressions, the group of portfolio-held loans is omitted from the estimation, and thus assumed to be the reference group. We cluster the standard errors at the zip code level to account for the fact that loans in the same geographical area are likely to suffer correlated (unobserved) shocks.

According to the estimates in the first column, private-label loans are about 0.3% less likely to receive concessionary modifications compared with loans held in portfolio. This estimate is economically small but statistically significant at the 10% level. When we consider all modifications the point estimate is 0.2% (statistically insignificant), and for the third specification, private-label loans actually are actually 0.9% more likely to receive concessionary modifications (statistically significant). As discussed above, all of these specifications include a number of additional loan characteristics that are important in the underwriting process, and thus, likely to play an important role in the modification decision. The first observation to make regarding the results reported in Panel B is that the difference between the incidence of modification for portfolio-held loans and private label loans becomes even smaller when these variables are controlled for in the estimation. The results also imply that

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<sup>21</sup>Equity and periods of unemployment are very important determinants of a borrower’s decision to default, and thus should also be important factors in the modification decision.

loans with higher credit scores are modified less, loans with higher ltv ratios are modified less, larger loans are modified more, and loans with more equity at the time of delinquency are modified less. We find a sizeable difference in terms of the frequency of modification for both refinances and for subprime loans. Conditional on being 60-days delinquent, subprime loans are modified about 2% more than prime loans. We estimate a model separately for subprime loans in Table 6.

Censoring is an important issue for any sample of mortgages, as there are currently many delinquent loans that are, or will, soon be good candidates for modification, as the housing market continues to decline. For this reason, we estimate a Cox proportional hazard model of the transition from serious delinquency to modification. The Cox model is very common in the survival analysis literature, and has the advantage of being both flexible in terms of functional form considerations, as the baseline hazard function can be treated as an incidental parameter, and easy to estimate in terms of computational considerations. The results, expressed as hazard ratios, are reported in Panel C. A hazard ratio less than 1 indicates that private-label loans are less likely to receive a modification compared to portfolio loans, while a ratio greater than 1 signifies the opposite. The estimates are consistent with what we report for the logits in the previous panel. Private-label loans are less likely to receive concessionary modifications, but this coefficient is statistically insignificant. For the our other two modification definitions the sign flips, but again the result is not statistically significant. All three specifications include the same covariates as the logit models.

## 3.2 Subsample Results

Table 6 contains further logit estimation results for various subsamples of interest to see if there are different effects compared to the full sample. Since the subprime indicator seems to be such a powerful predictor of modification conditional on serious delinquency in Table 5, we report the estimated marginal effects for only the sample of subprime loans in the second column of Table 6. In the third column, we report results from the sample of LPS mortgages for which the borrower had a FICO score of less than 620. This FICO cutoff is commonly used in the literature as an alternative definition of a subprime mortgage, and provides a nice robustness check. In the fourth column and fifth columns, we focus on samples of loans that we believe to contain the most amount of information regarding the borrowers, in order to try to minimize the amount of unobservables that could be potentially biasing the results. In the fourth column we focus on the sample of loans for which both the debt-to-income (DTI) ratio and the documentation status contains non-missing values,

while the fifth column contains results for only the loans that were fully documented (in terms of income and assets) at origination. Panel A contains both unconditional means and estimated marginal effects for concessionary modifications, while Panel B contains results for the broader definition that also includes non-concessionary modifications.

The results are largely consistent with those contained in Table 5. We re-display the results from the full sample in the first column of Table 6 for ease of comparison. The difference in modification frequency between private-label and portfolio-held, subprime mortgages for 60-day delinquent loans is small, and not statistically different from zero for both definitions of modification. Using a FICO cutoff of 620 as an alternative definition of subprime, does not seem to make too much of a difference. The unconditional means are smaller (for both types of loans) compared to the LPS subprime sample, as the LPS definition includes most of the loans with a FICO less than 620, but also some loans with higher associated FICOs. However, the marginal effects of private-label estimated from the logit models are quite similar to those from the LPS subprime sample, as they are economically small, and not statistically significant from zero. Finally, we also find small and largely insignificant results for the last two subsamples, displayed in the fourth and fifth columns of Table 6. Although, it is worth pointing out that we do find a statistically significant, positive estimate of private-label loans for the broad definition of modification (Panel B).

### 3.3 Alternative Delinquency Definition

As an additional robustness check, we broaden our definition of delinquency, and focus on modifications performed on loans subsequent to their first 30-day delinquency, which corresponds to one missed mortgage payment. While waiting until a borrower becomes seriously delinquent (defined as 60-days) to renegotiate is common practice in the servicing industry, there are no direct contractual stipulations (to our knowledge) that restrict a servicer from modifying a borrower who is 30-days delinquent. Thus, in Table 7 we repeat our analysis in Tables 5 and 6, but condition on 30-days delinquency rather than 60-days. The table contains three panels of estimation results, one for each of our modification definitions, and all of the subsamples described considered in Table 6. The unconditional means, logit marginal effects, and Cox hazard ratios are all reported for each combination of subsample and modification definition.

The results are very similar to those from the analysis of 60-day delinquent loans. According to the full sample and subprime sample logit models, portfolio loans receive slightly more concessionary modifications, and the differences (0.3 and 0.5 percentage points respectively) are statistically significant at conventional levels. However, according to the

subprime sample and full documentation sample Cox models, private-label loans actually receive more concessionary modifications, although those differences are also small.<sup>22</sup> The results for our second modification definition are similar, although we find more evidence of statistically significant, positive differences between the incidence of portfolio and private-level modifications. The sample of portfolio loans with non-missing information for DTI and documentation status are modified more often compared to the corresponding sample of private-label loans, but the magnitudes are still relatively small (10 to 20 percent difference relative to the unconditional mean). Finally, in Panel C, we see strong evidence for both the logit and Cox specifications, that delinquent private-label loans prepay more often than portfolio loans. The differences are statistically significant for every one of the subsamples.

### 3.4 Re-default Probabilities and Cure Rates

In the previous subsections, we showed that there is little difference in the frequency to modify mortgages between servicers of loans held in a private trust and loans held in portfolio. There are two potential reasons for why those exercises may fail to pick up important differences in servicer behavior that may truly exist. Second, it may be the case that contract frictions in securitization trusts do not result in substantial differences in the frequency of modifications, the extensive margin, but do result in significant differences in the intensive margin, with respect to the types of modifications performed, the extent to which contract terms are modified, and more broadly the care or effort which is expended in each modification by private-label servicers compared to portfolio servicers. Second, there may be a type of renegotiation that our algorithm does not identify, but that is used to a large extent in loss mitigation efforts, and used differently by servicers of private-label loans compared to servicers of portfolio loans. For example, forms of forbearance, which are often called re-payment plans in the industry, would not be picked up by our algorithm.<sup>23</sup> In this subsection, we use the LPS data to attempt to address these possibilities.

We perform two separate empirical exercises to address each these concerns in turn. First, we compare re-default rates of private-label modified loans with those of portfolio modified loans. We define re-default as a loan that is 60 days delinquent or more, in foreclosure process or REO 6 months after the time of the modification. If there are important

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<sup>22</sup>The logit marginal effects correspond to percentage point differences, while the Cox hazard ratios correspond to percent differences. If one expresses the logit marginal effects as a percent change of the unconditional means, those percent changes are very similar in magnitude to the Cox results.

<sup>23</sup>However, as we argued above, PSAs do not contain restrictions on re-payment plans, because they do not involve changing the terms of the mortgage. Thus, we would argue that differences in forbearance behavior that might exist, could not be the result of contract frictions in securitization trusts.

differences in the manner by which servicers of private-label loans modify mortgages relative to servicers of portfolio loans, then we would expect to see significant differences in the subsequent performance of modified loans.

Second, to address the possibility that our algorithm misses an important aspect of renegotiation, we compare the cure rates of seriously delinquent, private-label loans to those of seriously delinquent portfolio loans. The idea behind this exercise, is that any appreciable difference in servicer renegotiation behavior, will manifest itself in differences in cure rates. It is important to stress however, that differences in servicer renegotiation behavior are only one potential explanation for differences that may exist in cure rates. To put this idea in the terms of logical reasoning, differences in cure rates are a necessary condition for differences in renegotiation behavior, but are not a sufficient condition.

Table 8 contains the results from the re-default analysis. The first observation to note from the table is that the unconditional probability that a modified mortgage redefaults in this six month period is very large, at about 20-40 percent for payment reducing modifications (Panel A), and 40-50 percent for all modifications (Panel B). We will argue below, that the high level of re-default rates could explain why we observe so few modifications - very often they do not lead to successful outcomes even as little as 6 months after the modification. The second observation to note is that there is no statistically significant difference between the re-default rates of private-label loans relative to portfolio loans, once the observable characteristics of the mortgages are taken into account (this is valid for all of the subsamples). These results, combined with the statistics displayed in Table 4 suggests that there are not substantial differences in either the type of modification employed or in the care/effort expended by the two types of servicers.

Table 9 shows the results of logit models for the probability that a seriously delinquent loan subsequently cures. Our definition of a cure is that the loan is either current, 30-days delinquent, or prepaid after 12 months following the first 60-day delinquency. The first important point to make is that the unconditional cure probabilities are large (around 30 percent). Given that the unconditional modification probability is about 8 percent, this means that many loans cure without any intervention on the part of servicers. The second important observation to note in this table, is that the cure probabilities for portfolio loans and private label loans are quite similar. The unconditional cure probability is smaller by about 4.4 percentage points for private-label loans in the whole sample, but that is reduced to only 2.2 percentage points (statistically significant) when we control for observable characteristics of the loans and borrowers. We also include results for the subsamples of interest in columns 2-5. For each of the subsamples the sign of the difference actually reverses, as private-label loans are *more* likely to cure (the marginal effects are statistically significant,

with the exception of the  $FICO < 620$  sample). This is an important robustness check, as we argued above that unobserved heterogeneity is likely to be less of a problem in the subsamples (especially for the non-missing documentation status and DTI ratios sample and the full documentation sample). Thus, the change in the sign of the differences in cure rates between private-label servicers and portfolio servicers suggests that unobserved heterogeneity between the two loan types plays an important role.

## 4 Understanding the Empirical Results

If securitization does not block renegotiation, then why is it so rare? In this section, we build a simple model of the renegotiation decision which, in a stylized way, mirrors the net present value (NPV) calculation that servicers are supposed to do when deciding whether to offer a borrower a modification. We show that servicer uncertainty about whether the borrower will re-default even after successful renegotiation or uncertainty about whether the borrower will cure without renegotiation can dramatically affect the NPV calculation, and ruin what a naive observer might think of as a “win-win” deal for the borrower and lender. While many proponents of modification are aware of the former problem, “redefault risk,” none seem to be aware of the latter problem, what we will call “self-cure risk.”

In addition to the model, we also provide institutional evidence in this section that supports our arguments and findings above. This includes evidence of low modification frequencies in previous housing busts, well before the advent of securitization trusts; the equal treatment provision statements contained in the PSAs, which direct the servicer to behave as if it was in fact the investor of the mortgage-backed security, and thus the owner of the mortgages; and finally the absence of lawsuits to-date directed at servicers by investors in mortgage-backed securities, which one would expect to find if modifications are unambiguously better than foreclosures from an NPV calculation.

### 4.1 A Simple Model of Loss Mitigation

We consider a simple model of a lender’s decision to modify a delinquent loan.<sup>24</sup> There are three periods:  $t = 0, 1, 2$ . The borrower owes a mortgage payment of size  $m$  at time 1, and is due to repay the loan balance  $M$  in period 2. The mortgage is collateralized by a house, which is worth  $P_1$  and  $P_2$  in periods 1 and 2 respectively. In period 0, the lender has to make a decision to either modify the loan, or do nothing. If the lender fails to modify the loan,

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<sup>24</sup>Our model shares some basic similarities with the approach in Ambrose and Capone (1996), who also identify a role for self-cure risk in assessing the profitability of a loss mitigation action.

then with probability  $\alpha_0$ , the borrower will default in period 1, and the lender will foreclose and recover  $P_1 - \lambda$ , where  $\lambda$  is the cost of foreclosing on the property. If the borrower does not default next period, then the lender receives the periodic payment  $m$  in period 1, and the borrower repays the loan in full in period 2. The value to the lender of the loan without modification equals the present discounted value of the cash flow:

$$\alpha_0 * \max[(P_1 - \lambda), M] + (1 - \alpha_0)[m + (1/R)M]. \quad (1)$$

where we ignore discounting for the first period because there is no income in period 0. If the lender modifies the loan, then we assume that the borrower makes a reduced periodic payment  $m^*$  in period 1 with certainty, but then either defaults with probability  $\alpha_1$ , or repays a modified amount  $M^*$  in period 2. The value to the lender of the modified loan is:

$$m^* + (1/R)\alpha_1 * \max[(P_2 - \lambda), M^*] + (1 - \alpha_1)(1/R)M^*. \quad (2)$$

Taking the difference between expression (2) and (1) yields the following proposition.

**Proposition 1** *Modification makes sense if:*

$$\begin{aligned} &(\alpha_0 - \alpha_1)[m^* + \frac{1}{R}M^* - \max[(P_1 - \lambda), M]] \\ &\quad - (1 - \alpha_0)[m + \frac{1}{R}M - (m^* + \frac{1}{R}M^*)] \\ &\quad + \alpha_1[m^* + \frac{1}{R}\max[(P_2 - \lambda), M^*] - \max[(P_1 - \lambda), M]] > 0 \end{aligned} \quad (3)$$

To interpret equation (3), divide the population of borrowers into three groups. The first group, with mass of  $\alpha_0 - \alpha_1$  are borrowers who will repay in full with a modification but who will default otherwise. For this group, the investor gains the difference between the present value of the modified repayment  $m^* + \frac{1}{R}M^*$  and the recovery given foreclosure,  $\max[(P_1 - \lambda), M]$ . The second group, with mass  $1 - \alpha_0$ , includes borrowers who will repay whether or not they receive a modification. For this group, the investor loses the difference between full repayment and modified repayment. Gerardi and Willen (2009) refer to the first two terms as Type I error and Type II error respectively, in analogy with the statistical concepts. In this context, Type I error corresponds to the cost of not renegotiating loans that need modifying, while Type II error corresponds to the cost of modifying loans that will be repaid in the absence of assistance. The third term, with mass  $\alpha_1$ , includes borrowers who will default regardless of whether they receive a modification. For these borrowers, modification yields a periodic payment, but postpones foreclosure. Whether this is good or bad for the lender depends on the evolution of house prices, and the rate at which the

lender discounts the cash flow.

To illustrate the implications of the model, we perform some simple comparative statics. All else equal, an increase in  $\alpha_0$  makes modification more attractive to the investor, while an increase in  $\alpha_1$  makes modification less attractive. Intuitively, a higher  $\alpha_0$  means higher Type I error and lower Type II error, and a higher  $\alpha_1$  implies higher Type II error. Since in general, one would think that  $\alpha_0$  and  $\alpha_1$  would move in the same direction across borrowers, it is useful to note that an increase the gap,  $\alpha_0 - \alpha_1$ , makes modification more attractive.

We make three points about the model. First, when looking at the data, it is not sufficient to show that you would recover more from a modified loan than from foreclosure *ex-post*, to prove that modification is *ex-ante* optimal. To prove that a modification makes sense from the perspective of the lender, you must show that the Type I error, the modified loans that would have defaulted, exceeds the Type II error, the modified loans that would have paid off in the absence of modification. White (2009) among many others focuses entirely on Type I error:

The average loss for the 21,000 first mortgages liquidated in November was \$145,000, representing an average loss of 55% of the amount due. Losses on second lien mortgages were close to 100%. In comparison, for the modified loans with some amount of principal or interest written off, the average loss recognized was \$23,610. This seven-to-one difference between foreclosure losses and modification write-offs is striking, and lies at the heart of the failure of the voluntary mortgage modification program. At a minimum, there is room for servicers to be more generous in writing down debt for the loans they are modifying, while still recovering far more than from foreclosures in the depressed real estate market of late 2008. I will consider some of the reasons for this apparently irrational behavior in a later section.<sup>25</sup>

To see why this is wrong, take an extreme example with  $\alpha_1 = 0$ . In that case, the gain to modifications equals

$$\alpha_0[m^* + \frac{1}{R}M^* - \max[(P_1 - \lambda), M]] - (1 - \alpha_0)[m + \frac{1}{R}M - (m^* + \frac{1}{R}M^*)] \quad (4)$$

With  $\alpha_0$  sufficiently low, modification will not make sense. To be clear, our criticism of White (2009) and others has nothing to do with the possibility that the modified loan will default, as we have assumed here that the modified loan will pay off in full.

The second point here is that both the rate at which lenders discount future payoffs, and the evolution of prices affects the gains to modification. For mass  $(1 - \alpha_1)$  of the

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<sup>25</sup>White (2009), p. 14-15

borrowers, modification will simply delay foreclosure. In that case, the lender will get some extra income from any mortgage payments the borrower makes before re-defaulting, but the lender has to wait longer to obtain the final payout, and will get less if prices fall continue to fall.

The third point is that the lender’s information set plays a crucial role here, and one could argue that it should only contain information outside the control of the borrower. This would limit the set to the origination characteristics of the loan, prices, and interest rates. Employment status, income, and marital status all present problems, although they can be partially overcome – as in the case of unemployment insurance. Delinquency status, which seems like a natural candidate, is a difficult issue. On one hand, a borrower has virtually complete control over it. On the other hand, it is a costly signal, as a 60-day delinquency does adversely affect one’s credit history and future access to credit markets. Thus, when considering ways to design a profitable modification program, which implies trying to maximize  $\alpha_0$  and minimize  $\alpha_1$ , a lender must restrict its information set to a relatively small set of variables that are contemporaneously exogenous to the borrower.

## 4.2 Institutional Evidence

While the results from section 3 may be surprising to market commentators who believe that contract frictions inherent in securitization trusts are preventing large-scale modification efforts in mortgage markets, we argue in this section that both historical evidence, and evidence from securitization contracts, actually support our findings.

First, we look at history. If securitization, or more precisely private-label securitization, inhibits renegotiation, then we would expect that renegotiation would have been common in the 1990s, when there was little private-label securitization or in the 1970s, when securitization itself was rare. But what historical evidence we have does not bear that out. In 1975, Touche Ross surveyed loss mitigation activities at savings and loans and found that, “Lenders... were unwilling to either modify loans through extended terms or refinancing to a lower rate,”<sup>26</sup>. In the 1990s, a report commissioned by Congress to study foreclosure alternatives, said that, “Along with loan modifications, long-term forbearance/repayment plans are the most underutilized foreclosure avoidance tools currently available to the industry.”<sup>27</sup>

Second, many observers have focused on institutional factors that inhibit loan modification when the loan is securitized, but other factors may play a similar role for portfolio lenders as well. In particular, accounting rules force lenders to take write-downs at the time

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<sup>26</sup>Capone (1996), p. 20-21.

<sup>27</sup>Capone (1996), p. x.

of the modification, to identify modified loans as troubled debt restructurings and also impose burdensome reporting requirements for modified loans, all of which make modification costly to a bank. Downey Financial, for example, attempted to refinance current borrowers out of risky Option ARMs into safer fixed rate instruments and argued that the change should not affect their balance sheet because the borrowers had never missed payments, but their accountants viewed the refinancings as “troubled debt restructurings” and forced them to restate the share of nonperforming assets for November 2007 to 5.77 percent from 3.65 percent.<sup>28</sup>

If modifications were truly in the best financial interest of MBS investors, as many commentators have alleged, we would expect to see concern on their part regarding the low levels of modifications performed to-date. But, according to Cordell, Dynan, Lehnert, Liang, and Mauskopf (2008b), who interviewed a number of MBS investors, they (the investors) are not concerned that servicers are foreclosing on many more mortgages than they are modifying. Thus, there does not seem to be much concern by market participants that either incentives or contract frictions are inhibiting servicers from performing loan modifications. The evidence in the literature seems to suggest a small role for contract frictions in the context of renegotiation. In a 2007 study of a small sample of PSAs, Credit Suisse found that less than 10 percent of the contracts ruled out modifications completely, while approximately 40 percent allowed modifications, but with quantity restrictions<sup>29</sup>, and the rest, about half, contained no restrictions on renegotiation behavior. Hunt (2009) also analyzed a sample of subprime PSAs, and concluded that outright modification bans were extremely rare. A 2008 report by the Congressional Oversight Panel (COP) for the Troubled Asset Recovery Program (TARP), analyzed a number of securitized mortgage pools with quantity restrictions, and concluded that none of the restrictions was binding. In terms of incentive issues, Hunt (2009) found that most of the contracts in his sample explicitly instructed the mortgage servicer to behave as if it were the owner of the pool of the loans:

The most common rules [in making modifications] are that the servicer must follow generally applicable servicing standards, service the loans in the interest of the certificate holders and/or the trust, and service the loans as it would service loans held for its own portfolio. Notably, these conditions taken together can be read as attempting to cause the loans to be serviced as if they had not been securitized. (p. 8, insertion added)

In addition to the evidence from securitization contracts, evidence on the frequency of

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<sup>28</sup><http://www.housingwire.com/2008/01/14/downey-financial-accounting-rules-suck/>.

<sup>29</sup>The quantity restrictions often took the form of a limit (usually 5 percent) on the percentage of mortgages in the pool that could be modified without requesting permission from the trustee.

modifications in previous recessions and housing busts also supports the idea that securitization is not an important force holding up renegotiation between borrowers and lenders in mortgage markets.

## 5 Conclusion

There is widespread concern that an inefficiently low number of mortgages have been modified during the current crisis, which leads to excessive foreclosure levels and leaves both families and investors worse off. We use data on loans originated between 2005 and 2007 to shed more light on the determinants of mortgage modification, with a special focus on the claim that delinquent loans have different probabilities of renegotiation depending on whether they are securitized or held in a bank's portfolio. By comparing the relative frequency of modification between securitized and non-securitized mortgages we can shed light on the question of whether institutional frictions in the secondary mortgage market are inhibiting the modification process from taking place.

We find that unconditionally, servicers actually modify *more* securitized mortgages but this results largely from the fact that securitized mortgages are more risky and when we control for the characteristics of the loan and the borrower, this gap vanishes. Our results on the role of securitization in inhibiting effect modification are more ambiguous: Modifications of securitized loans appear to be less effective than modifications of portfolio loans. Our best guess is that the foreclosure hazard of modified mortgages is between 10% and 25% higher than for loans held in the servicer's portfolio. While this may seem large, we make two caveats. First, the results are very sensitive to the specification and the sample. For at least some versions of the model, securitized loans actually default less. Second, because so few delinquent loans receive modifications, the overall effect on foreclosures is small.

## 6 Appendix A – Identifying Modifications in the LPS Dataset

In this section we discuss in detail the assumptions that we used to identify modified loans in the LPS dataset. The LPS dataset is updated on a monthly basis, and the updated data include both new mortgages originated, as well as a snapshot of the current terms and delinquency status of outstanding mortgages. Essentially, for a given mortgage, we compare the updated terms to the terms at origination, as well as the change in terms from the proceeding month, and if there is a material change over-and-above the changes stipulated in the mortgage contract, then we assume that the contract terms of the mortgage have been modified.

### 6.1 Interest Rate Reductions

We use a different set of rules to identify reduced interest rates for fixed-rate mortgages (FRM) and adjustable-rate mortgages (ARM). In principle, identifying a rate change for an FRM should be easy, since by definition the rate is fixed for the term of the mortgage. However, after a detailed inspection of the LPS data, it became apparent that some of the smaller rate fluctuations were likely due to measurement error rather than to an explicit modification. Thus, we adopt a slightly more complex criteria: The difference between the rate at origination and the current rate must be greater than 50 basis points; *and* the difference between the rate in the previous month and the current rate must be greater than 50 basis points; *and* either the mortgage must be 30-days delinquent, the loans is currently in loss mitigation proceedings (as reported by the servicer), or the difference between the rate in the previous month and the current rate must be greater than 300 basis points (which allows for the possibility that a loan that is current could feasibly qualify for a modification).

Identifying interest rate reductions for ARMs is slightly more complicated since by definition the interest rate is variable and can move both up and down. The LPS data contain the information necessary to figure out how much the interest rate should move from month to month. This rate is often referred to as the fully-indexed rate, as it is normally specified to be a fixed spread above a common nominal interest rate. The LPS contains information regarding the initial rate, the appropriate index rate, and the spread between the index and the mortgage rate. In addition, the majority of ARMs are characterized by a period at the beginning of the contract in which the interest rate is held constant (these mortgages are often referred to as hybrid ARMs). At the end of this period, the interest rate adjusts (or resets) to a certain spread above an index rate, and then subsequently adjusts at a specific

frequency. The LPS dataset contains information regarding the length of the initial fixed period, and thus we can identify this period in the data and determine at the point at which the interest rate should begin to adjust (we refer to this period as the reset date). Our criteria for identifying an interest rate reduction for an ARM is as follows: The difference between the rate at origination and the current rate must be greater than 50 basis points; *and* the difference between the rate in the previous month and the current rate must be greater than 50 basis points; *and* if the reset date has passed then the difference between the fully-indexed rate and the current rate must be at least 100 basis points ; *and* either the mortgage must be 30-days delinquent, the loans is currently in loss mitigation proceedings (as reported by the servicer), or the difference between the rate in the previous month and the current rate must be greater than 300 basis points (which allows for the possibility that a loan that is current could feasibly qualify for a modification). In addition, we allow for more modest month-to-month decreases in the interest rate (200 to 300 basis points) as long as there is also a positive change in the delinquency status of the loan (i.e. the loan is reported to be less delinquent). Our inspection of the data suggests that the majority of modifications involve a resetting of the delinquency status back to current, or minor delinquency, so conditioning on this change likely eliminates many false positives.

## 6.2 Term Extensions

Term extensions, in theory, should be straightforward to identify in the LPS data, but due to possible measurement error in the variable that measures the remaining maturity of each loan. We defined a term extension in the LPS dataset to be a case in which the loan was at least 30-days delinquent at some point; and the remaining number of years left increases by at least 20 months *or* the change in number of years left is greater than the difference between the original term of the loan and the remaining term (for example, if the original maturity is 360 months, and the loan has 350 months remaining, then the increase in length must be at least 10 months); and finally, either the monthly payment must decrease, *or* the principal balance must increase, *or* the loan must be in loss mitigation.

## 6.3 Principal Balance Reductions

Reductions in the remaining balance of a mortgage is perhaps the most difficult modification to identify because of the prevalence of “curtailment” or partial prepayment among mortgage borrowers. For example, it is common for borrowers to submit extra mortgage payments in order to pay down the loan at a faster rate. For this reason, we were forced to adopt strict criteria to limit the number of false positives. Our criteria for identifying a principal

balance reduction is as follows: The month-to-month decrease in the remaining principal balance must be at least -10% and cannot be more than -30% (the upper bound does not matter as much as the lower bound – we experimented with -40% and -50%, but did not find a substantial difference); the principal balance recorded in the previous month must be greater than \$25,000 (since we throw second liens out, and are looking at only mortgages originated after 2004, this cutoff does not bind often); the month-to-month payment change must be negative (there are only a few cases in which the principal balance is reduced without corresponding decrease in the payment, but in these cases the term is extended, and thus is picked up in our code for identifying term extensions); and finally, the mortgage must be either 30-days delinquent or currently in loss mitigation proceedings (as reported by the servicer).

## 6.4 Principal Balance Increases

For interest-only and fully-amortizing mortgages, identifying an increase in the principal balance due to the addition of arrears is relatively straightforward. It becomes trickier for mortgages that allow for negative amortization, as the principal balance is allowed to increase over the course of the contract, by definition. For interest-only and fully-amortizing mortgages our criteria are: The month-to-month principal balance must increase by at least 0.5% (to rule out measurement error in the data); the loan must have been at least 30-days delinquent at the time of the balance increase; and finally, the month-to-month payment change must be positive unless there is also a corresponding increase in the term of the loan. For mortgages that allow for negative amortization, the criteria are similar except that the balance increase must be at least 1% and there must be a positive change in the delinquency status of the loan.

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Table 1: Examples of Modifications in the data.

**Example 1:** Servicer cuts interest rate, capitalizes arrears in the balance of the loan and extends term to 40 years.

Date	MBA Delinq. Stat.	Interest Rate	Monthly Payment	Outstanding Balance	Remaining Term in Months
2008m10	9	6.5	907	141,323	340
2008m11	9	6.5	907	141,323	339
2008m12	9	6.5	907	141,323	338
2009m1	C	4.5	660	146,686	479

**Example 2:** Servicer capitalizes arrears into the balance of the loan but otherwise leaves the loan unchanged.

Date	MBA Delinq. Stat.	Interest Rate	Monthly Payment	Outstanding Balance	Remaining Term in Months
2008m5	6	9.25	1,726	208,192	346
2008m6	9	9.25	1,726	208,192	346
2008m7	9	9.25	1,726	208,192	346
2008m8	C	9.25	1,815	218,316	341
2008m9	C	9.25	1,815	218,184	340

Table 2: Robustness of the modifications algorithm

**False Positives by type of modifications**

	# of Modifications Using WF CTS Data	False Positives
FRM Rate Reduction	5,381	8.0%
ARM Rate Reduction	8,951	22.0%
Principal Reductions	470	1.9%
Principal Increases	13,010	12.8%
Term Increases	394	2.3%

**Overall success of algorithm**

	No Mod Using Our Algorithm	Mod Using Our Algorithm	Total
No Mod in WF Data	2,329,187	3,559	2,332,746
Mod in WF Data	3,627	17,514	21,141
Total	2,332,814	21,073	2,353,887

Notes: We test our algorithm on a dataset of securitized mortgages in which the trustee has identified modifications (data is from Wells Fargo Trustee Services). The lower panel shows that about 17.2% of our modifications are false positives, meaning that we identify modifications but the trustee does not and about 16.9% are false negatives, meaning that the trustee identifies a modification but we do not.

Table 3: Modification Statistics

**(1) By Type of Modification: 2007:Q1–2008:Q4**

	# Loans Modified	Interest Rate Reductions		Principal Balance Reductions		Principal Balance Increases		Term Extensions	
		#	(% total)	#	(% total)	#	(% total)	#	(% total)
2007:Q1	10,940	600	5.3	700	6.2	8,660	76.4	1,380	12.2
2007:Q2	14,600	820	5.4	550	3.7	11,630	77.3	2,050	13.6
2007:Q3	17,720	770	4.1	810	4.3	15,170	81.2	1,940	10.4
2007:Q4	27,150	2,990	9.7	700	2.3	22,520	72.8	4,740	15.3
2008:Q1	36,230	6,010	13.8	900	2.1	32,100	73.8	4,500	10.3
2008:Q2	44,750	9,050	16.4	1,300	2.4	39,750	72.1	5,030	9.1
2008:Q3	62,190	16,280	20.3	940	1.2	56,940	70.9	6,110	7.6
2008:Q4	74,800	28,630	26.7	1,450	1.4	65,960	61.5	11,230	10.5

Notes: These statistics were computed using a 10% random sample of the LPS data. Quantities obtained from the data are multiplied by a factor of 10. The percentages are taken with respect to the total number of modifications, and *not* loans modified. Thus, there is double-counting in the sense that some loans received multiple types of modifications in a given quarter.

**(2) By Payment Change**

	Payment Decreases					Payment Increases				
	#	mean $\Delta$		median $\Delta$		#	mean $\Delta$		median $\Delta$	
		\$	%	\$	%		\$	%	\$	%
2007:Q1	2,080	-492	-13.2	-157	-10.0	5,020	106	6.7	62	4.4
2007:Q2	2,060	-464	-12.7	-141	-9.6	7,710	120	7.0	63	4.4
2007:Q3	2,470	-290	-12.9	-125	-9.7	10,380	110	6.7	60	4.3
2007:Q4	5,600	-367	-15.3	-159	-11.7	14,540	100	5.9	59	3.9
2008:Q1	11,500	-358	-14.0	-210	-13.2	18,720	108	6.5	62	4.3
2008:Q2	18,660	-425	-16.1	-239	-14.1	20,770	124	7.4	69	4.1
2008:Q3	31,770	-562	-21.5	-365	-20.2	26,400	124	6.3	63	3.6
2008:Q4	48,000	-503	-22.9	-315	-21.7	22,520	104	6.0	53	3.6

Notes: These statistics were computed using a 10% random sample of the LPS data. Quantities obtained from the data are multiplied by a factor of 10.

**(3) Loan Characteristics of Modified Mortgages**

	All Loans					Modifications				
	#	mean	p25	p50	p75	#	mean	p25	p50	p75
FICO (at origination)	1,892,777	706	660	713	762	17,533	622	580	621	662
LTV (at origination)	2,250,162	75	67	79	85	21,675	82	78	80	90
DTI (at origination)	1,346,093	37	28	38	45	13,945	41	35	41	47
Mortgage balance (at origination)	2,267,497	231K	121K	185K	288K	21K	234K	121K	186K	294K
<i>% characterized as</i>										
LTV = 80			14.4					21.7		
Subprime			6.8					47.4		
Fixed			71.2					39.7		
Hybrid ARM			7.7					26.2		
IO-ARM			11.3					13.1		
IO-Fixed			2.1					2.7		
Option-ARM			5.1					12.0		
Option-Fixed			0.3					1.4		
Owner			89.3					96.0		
Investor			7.1					2.6		
Vacation Home			3.7					1.1		
Purchase			51.9					49.0		
Low/no documentation			29.2					20.4		

Notes: These statistics were computed using a 10% random sample of the LPS data. The sample includes mortgages originated after 2004.

Table 4: Modification Comparison by Payment Change

<i>Private Label Modifications</i>										
	Payment Decreases					Payment Increases				
	#	mean		median		#	mean		median	
		\$	%	\$	%		\$	%	\$	%
2007:Q1	106	-614	-14.42	-162	-10.85	239	121	6.02	76	3.37
2007:Q2	110	-505	-12.02	-222	-9.30	364	168	7.96	76	3.49
2007:Q3	128	-261	-11.82	-131	-8.42	558	145	7.52	75	3.65
2007:Q4	288	-313	-13.38	-163	-12.36	741	125	6.24	74	3.52
2008:Q1	634	-393	-16.12	-261	-15.65	938	133	6.76	79	4.08
2008:Q2	1,014	-540	-18.94	-334	-17.89	1,241	152	8.14	83	4.08
2008:Q3	1,778	-641	-22.01	-423	-19.95	1,805	137	6.22	70	3.31
2008:Q4	1,993	-565	-21.73	-367	-20.13	1,398	118	5.91	61	3.23

*Portfolio Modifications*

	Payment Decreases					Payment Increases				
	#	mean		median		#	mean		median	
		\$	%	\$	%		\$	%	\$	%
2007:Q1	28	-759	-20.90	-428	-17.19	128	106	7.78	52	5.46
2007:Q2	19	-1172	-25.17	-656	-28.07	222	81	6.11	55	5.28
2007:Q3	31	-395	-17.13	-168	-15.29	255	71	6.13	43	5.37
2007:Q4	90	-474	-11.11	-90	-2.48	292	70	5.50	37	4.29
2008:Q1	187	-369	-10.00	-183	-8.08	331	80	6.59	33	3.97
2008:Q2	309	-304	-10.90	-117	-6.64	405	63	5.59	34	3.56
2008:Q3	376	-585	-25.19	-295	-17.85	359	105	7.04	39	4.26
2008:Q4	616	-794	-31.91	-384	-25.04	389	59	5.48	35	3.51

Table 5: Modifications (Main Sample)

<b>Panel A: Unconditional Percentages</b>			
	Concessionary Mods	All Mods	All Mods + Prepayments
Portfolio	0.032	0.087	0.147
Private-label	0.026	0.084	0.155

<b>Panel B: Logit Regressions (12 month horizon)</b>			
	Concessionary Mods	All Mods	All Mods + Prepayments
Private-label	-0.003	0.002	0.009
	-1.69	0.58	1.95
Initial Rate	0.001	-0.004	-0.007
	1.45	-5.7	-7.25
<i>LTV</i> Ratio	0	0	-0.002
	-0.24	-1.68	-11.14
<i>LTV</i> = 80	0	-0.014	-0.034
	-0.18	-6.25	-11.7
<i>FICO</i>	0	0	-0.002
	-0.02	-0.43	-4.62
<i>FICO</i> <sup>2</sup>	0	0	0
	-0.39	-0.08	3.95
<i>FICO</i> < 620	0.002	0.029	0.034
	0.53	3.43	3.42
620 ≤ <i>FICO</i> < 680	0.005	0.017	0.024
	1.46	2.95	3.41
Log Original Amount	0.004	0.007	0.022
	3.12	2.96	7.47
Equity at delinquency	-0.001	-0.003	0
	-0.4	-1.09	0
Negative Equity	-0.006	-0.022	-0.022
	-1.6	-3.17	-1.77
Unemployment	0	-0.002	-0.005
	-0.37	-3.13	-4.37
Refi	0.006	0.015	0.04
	4.14	5.98	11.67
Subprime	0.02	0.037	0.042
	9.32	11.71	10.87
Other Controls	Y	Y	Y
# Mortgages	66,541	66,541	66,541

**Panel C: Duration Model**

	Concessionary Mods	All Mods	All Mods + Prepayments
Private-label	0.921	1.002	1.018
	-1.41	0.07	0.68
# Mortgages	87,343	87,343	87,343

Notes: Other controls include indicator variables for Jumbo, Option, Hybrid and Interest-Only mortgages, as well as for condos and multifamily homes. Panel B shows the marginal effects of logit regressions with a 12-month horizon, t statistics shown below the coefficients. Standard errors are clustered at the zip code level. Panel C shows hazard ratio estimates from a Cox proportional hazards model.

Table 6: Modifications (Robustness tests with alternative samples)

<b>Panel A: Concessionary Modifications</b>					
	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.032	0.047	0.034	0.028	0.023
Private-label Mean	0.026	0.037	0.031	0.033	0.037
Marginal Effect (private-label)	-0.003 -1.69	-0.004 -0.94	-0.003 -0.77	0 -0.14	0.007 1.46
# Mortgages	66,541	33,719	27,639	25,543	18,097

<b>Panel B: All Modifications</b>					
	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.087	0.111	0.097	0.092	0.077
Private-label Mean	0.084	0.103	0.109	0.107	0.124
Marginal Effect (private-label)	0.002 0.58	0.004 0.61	0.007 1.06	0.006 0.97	0.025 2.94
# Mortgages	66,541	33,719	27,639	25,543	18,097

Notes: Portfolio and Private label means are unconditional probabilities of modification in each sample. Marginal effects are computed from logit models with a 12-month horizon that include all the controls in Table 5. Standard errors are clustered at the zip code level. T-statistics are reported below the marginal effects.

Table 7: Modifications Conditional on 30 Days Delinquency (Logits)

<b>Panel A: Concessionary Mods</b>					
	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.014	0.025	0.016	0.014	0.012
Private-label Mean	0.014	0.021	0.016	0.017	0.019
Marginal Effect (Logit)	-0.003 -2.72	-0.005 -2.31	-0.001 -0.55	-0.002 -1.57	0.001 0.37
Hazard Ratio (Cox)	1.03 0.59	1.147 1.83	1.027 0.31	0.969 -0.42	1.237 2.34
# Mortgages	120,558	51,285	43,550	47,993	34,403

<b>Panel B: All Mods</b>					
	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.038	0.056	0.051	0.042	0.052
Private-label Mean	0.042	0.055	0.051	0.047	0.035
Marginal effect (Logit)	-0.004 -2.39	-0.007 -1.79	-0.004 -1.22	-0.008 -3.16	-0.001 -0.2
Hazard Ratio (Cox)	1.043 1.42	0.951 -1.05	1.008 0.17	0.909 -2.23	1.065 1.21
# Mortgages	120,558	51,285	43,550	47,993	34,403

<b>Panel C: All Mods + Prepayment</b>					
	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.145	0.195	0.152	0.147	0.13
Private-label Mean	0.174	0.211	0.218	0.185	0.198
Marginal effect (Logit)	0.023 7.31	0.021 2.98	0.044 6.46	0.016 3.47	0.029 4.54
Hazard Ratio (Cox)	1.158 9.09	1.05 1.69	1.181 5.72	1.098 3.88	1.202 6.56
# Mortgages	120,558	51,285	43,550	47,993	34,403

Notes: Portfolio and Private label means are unconditional probabilities of modification in each sample. Marginal effects are computed from logit models with a 12-month horizon that include all the controls in Table 5. Hazard ratios are computed from Cox proportional hazard models with the same controls as in Table 5. Z-statistics are shown below the coefficients, and T-statistics are reported below the marginal effects. Standard errors are clustered at the zip code level. Sample sizes refer to the logit regressions. The sample sizes for the Cox models are slightly larger.

Table 8: Re-default Conditional on Modification

Panel A: Payment Reducing Mods					
	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.308	0.386	0.332	0.228	0.249
Private-label Mean	0.358	0.392	0.371	0.362	0.359
Marginal effect (Logit)	0.016 0.66	-0.001 -0.03	-0.015 -0.35	0.03 0.81	-0.004 -0.1
# Mortgages	4,626	2,514	1,562	1,475	1,135

Panel B: All Mods					
	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.393	0.53	0.444	0.404	0.403
Private-label Mean	0.449	0.5	0.501	0.482	0.482
Marginal effect (Logit)	0.008 0.58	-0.023 -0.84	-0.009 -0.38	-0.021 -0.97	-0.033 -1.24
# Mortgages	14,796	7,073	5,344	4,594	3,620

Notes: Re-default is defined as loans that are 60 days delinquent, 90 days delinquent, in the process of foreclosure or in REO 6 months after the modification. Marginal Effects refer to the marginal effects of a logit model with a horizon of 6 months. T-stats shown below the marginal effects. Standard errors are clustered at the zip code level.

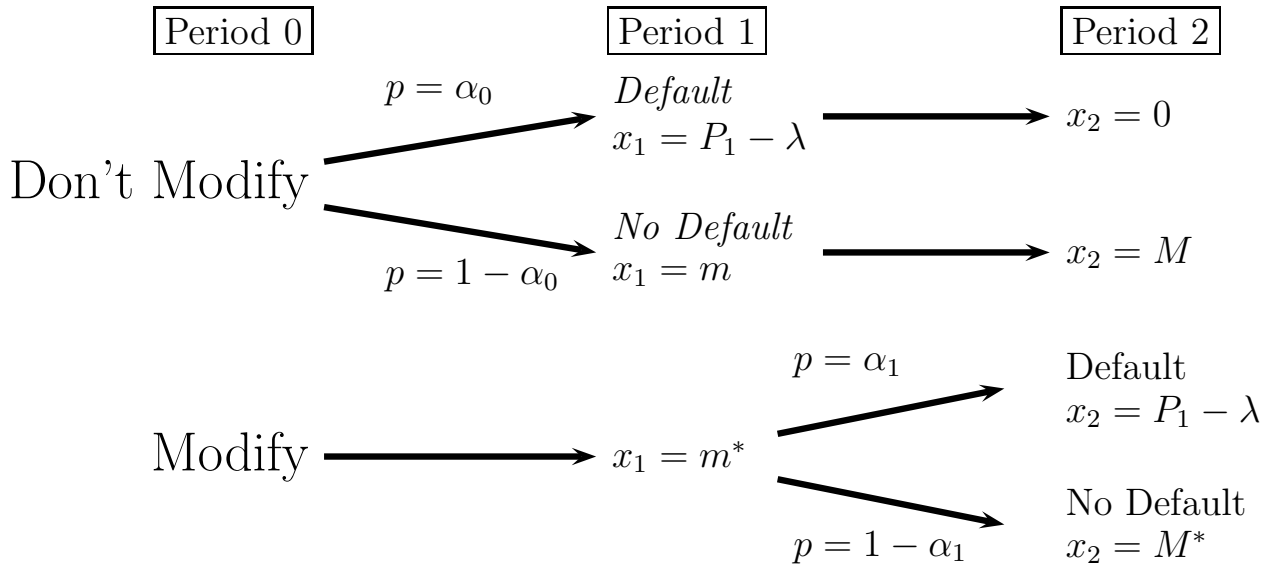
Table 9: Cure Conditional on 60 Days Delinquency

	All Loans	Subprime	<i>FICO</i> < 620	Non-missing Documentation and DTI	Fully Documented
Portfolio Mean	0.300	0.257	0.320	0.280	0.299
Private-label Mean	0.256	0.289	0.328	0.289	0.324
Marginal effect (Logit)	-0.022 -4.32	0.043 4.31	0.004 0.44	0.022 2.8	0.025 2.43
# Mortgages	66,451	33,719	27,639	25,543	18,097

Notes: The dependent variable (“Cure”) is defined as a loan that is either current, 30 days delinquent, or prepaid 12 months after the first 60-day delinquency. Portfolio and Private-label means are unconditional probabilities of modification in each sample. Marginal effects are computed from logit models with a 12-month horizon that include all the controls in Table 5. Standard errors are clustered at the zip code level. T-statistics are reported below the marginal effects.

Figure 1:

### (1) Model of loan modification



### (2) Understanding the gains to modification

Share of borrowers	$1 - \alpha_0$	$\alpha_0 - \alpha_1$	$\alpha_1$
Description	Borrower always repays Lender loses because borrower would have paid in full	Modification effective Lender gains because modified payments worth more than foreclosure	Borrower never repays Foreclosure is delayed May or may not help lender
Net gain	$m^* + \frac{1}{R}M^* - (m + \frac{1}{R}M)$	$m^* + \frac{1}{R}M^* - (P_1 - \lambda)$	$m^* + \frac{1}{R}(P_2 - \lambda) - (P_1 - \lambda)$
Error	“Type II error” Costly assistance to borrowers who can pay	“Type I error” Don't help borrowers who would have defaulted	“Type III error” Lender loses if $R$ is large or if $P_1 - P_2$ is big